

Peter C. B. Phillips' Contributions and Influence in Financial Econometrics

Jun Yu¹

¹Singapore Management University

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Outline

- 1 Overviews
- 2 Identification
- 3 Parametric Methods – Discretization Bias
- 4 Nonparametric Methods – Specification Bias
- 5 Finite Sample Issues – Estimation Bias
- 6 Other Contributions
- 7 References

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- 5 Nonlinear cointegration. References: Phillips and Park (2002). Financial Applications: Hu and Phillips (2004), Huang and Yu (2008).

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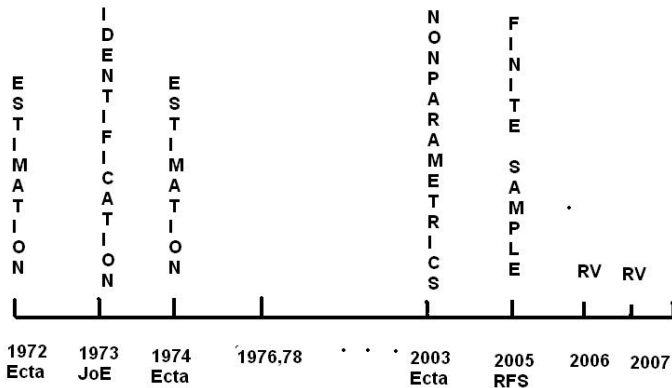
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Overview 3: Time Series Plot



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- The exact discrete time model is:

$$X_t = BX_{t-1} + \varepsilon_t, \quad (3)$$

where $B = \exp(Ah)$, $\varepsilon_t = \int_0^h \exp(sA)\xi(t-s)ds$.

The Identification Problem

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- Extensions: Hansen and Sargent (1983, Ecta).

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- Why exact discrete time model? Not subject to **discretization bias!**
- Euler discretization:

$$X_t = (1 - Ah)X_{t-1} + \epsilon_t.$$

Continue

- When $A = -1$, $h = 1/12$, e^{Ah} is 0.92 whereas $1 - Ah$ is 0.9167.
If $A = 1$, $h = 1$, then e^{Ah} is 0.3679 whereas $1 - Ah$ is 0.

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If $A = 1$, $h = 1$, then e^{Ah} is 0.3679 whereas $1 - Ah$ is 0.
- GLS procedure for the exact discrete time model:

$$\min(X_t - \exp(A(\theta)h)X_{t-1})'S(X_t - \exp(A(\theta)h)X_{t-1})$$

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- Gaussian estimation (QMLE).

The Extensions: Phillips (1974, Ecta)

- Extension 1 – *model with identities*

$$DX(t) = A(\theta)X(t) + H\xi(t),$$

where A is $n \times n$ but H is $n \times r$ with $r < n$.

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where $\varepsilon_t = \int_0^1 \exp(sA)H\xi(t-s)ds$.

- Conditions are needed to ensure the positivity of the covariance of ε_t .

- Extension 2 – *model with exogenous variables*

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where B is $n \times m$ and $Z(t)$ is a vector of exogenous variables.

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- Further extensions: mixed stock and flow variables, higher order models.

Interest Rate Models: Yu and Phillips (2001, Ects J)

- Models

$$dr(t) = \kappa(\mu - r(t))dt + \sigma(r_t)dB(t), \quad (4)$$

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$$r(t) = e^{-\kappa h}r(t-1) + \mu(1 - e^{-\kappa h}) + \int_0^h e^{-\kappa(h-\tau)}\sigma(r_{t+\tau})dB(\tau), \quad (5)$$

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- The error is not normal.

Continue

- To get an exact discrete model with Gaussian errors, Yu and Phillips (2001) used

$$r(t_{j+1}) = \mu(1 - e^{-\kappa h_{j+1}}) + e^{-\kappa h_{j+1}} r(t_j) + M(h_{j+1}), \quad (6)$$

where $M(h)$ is a continuous martingale with quadratic variation

$$[M]_h = \sigma^2 \int_0^h e^{-2\kappa(h-\tau)} \sigma^2(t + \tau) d\tau, \quad (7)$$

$$h_{j+1} = \inf\{s | [M_j]_s \geq a\} = \inf\{s | \int_0^s e^{-2\kappa(s-\tau)} \sigma^2(t_j + \tau) d\tau \geq a\}, \quad (8)$$

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- According to Dambis, Dambus-Schwarz theorem, $M(h_{j+1}) \sim N(0, a)$.

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$$E [dX(t) | X(t)] = \mu(X(t))dt, \quad (10)$$

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- Assume X_t is observed at equispaced points $(\Delta_{n,T}, 2\Delta_{n,T}, \dots, n\Delta_{n,T})$ over the period $[0, T]$. So $n = T\Delta_{n,T}$. It is assumed $n \rightarrow \infty$, $T \rightarrow \infty$ and $\Delta_{n,T} = T/n \rightarrow 0$.

Continue

- The nonparametric estimator of $\mu(X(t))$ is

$$\frac{\sum_{i=1}^n K\left(\frac{X_{i\Delta_{n,T}} - x}{h_{n,t}}\right) \tilde{\mu}_{n,T}(x)}{\sum_{i=1}^n K\left(\frac{X_{i\Delta_{n,T}} - x}{h_{n,t}}\right)}$$

- The nonparametric estimator of $\sigma^2(X(t))$ is

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$$\tilde{\mu}_{n,T}(x) = \frac{1}{m_{n,T}(i\Delta_{n,T})\Delta_{n,T}} \sum_{j=0}^{m_{n,T}(i\Delta_{n,T})-1} [X_{t(i\Delta_{n,T})j+\Delta_{n,T}} - X_{t(i\Delta_{n,T})j}]$$

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- Application 2: Corradi and Disto (2007)
- Application 3: Jeffrey, Linton, Nguyen, Phillips (2004)

Method 2: Bandi and Phillips (2006, JoE)

- Estimate parametric models via nonparametric methods

$$dX(t) = \mu(X(t); \theta_1)dt + \sigma(X(t); \theta_2)dB(t), \quad (12)$$

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- The estimator of θ_2 :

$$\min \|\tilde{\sigma}_{n,T}^2 - \sigma^2(X(t); \theta_2)\|$$

Method 3: Phillips and Yu, (JoE, forthcoming)

- Estimate parametric models via nonparametric and parametric methods

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Method 3: Phillips and Yu, (JoE, forthcoming)

- Estimate parametric models via nonparametric and parametric methods

$$dX(t) = \mu(X(t); \theta_1)dt + \sigma(X(t); \theta_2)dB(t),$$

- If the diffusion term is known (i.e. $\sigma(X_t; \theta_2) = \sigma(X_t)$), the exact continuous record log-likelihood is

$$\ell_{IF}(\theta_1) = \int_0^T \frac{\mu(X_t; \theta_1)}{\sigma^2(X_t)} dX_t - \frac{1}{2} \int_0^T \frac{\mu^2(X_t; \theta_1)}{\sigma^2(X_t)} dt.$$

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- On the other hand

$$[X]_T = \int_0^T (dX_t)^2 = \int_0^T \sigma^2(X_t; \theta_2) dt,$$

where $[X]_T$ is the quadratic variation of X , which can be consistently estimated by $[X_\Delta]_T$ as $\Delta \rightarrow 0$.

Method 4: Phillips and Yu (2006)

- Estimate integrated volatility using nonparametric methods

$$dp^*(t) = \sigma(t)dB(t), \quad (13)$$

where $p^*(t)$ is the equilibrium log-price and $\sigma(t)$ is a càdlàg volatility process. The quantity of interest is $IV = \int_0^1 \sigma^2(t)dt$.

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- Assume that $p^*(t)$ is observed at $\{t_{i,m} = \frac{i}{m} : i = 0, \dots, m\}$ with $h = \frac{1}{m}$

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- A nonparametric estimate is RV

$$\sum_{i=1}^m [p_{i,m}^* - p_{i-1,m}^*]^2 := RV^{(m)}(p^*).$$

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- Assume that $p^*(t)$ is observed at $\{t_{i,m} = \frac{i}{m} : i = 0, \dots, m\}$ with $h = \frac{1}{m}$
- A nonparametric estimate is RV

$$\sum_{i=1}^m [p_{i,m}^* - p_{i-1,m}^*]^2 := RV^{(m)}(p^*).$$

- Barndorff-Nielsen and Shephard (2002) showed that

$$\sqrt{m} \left[RV^{(m)}(p^*) - IV \right] | \sigma^2(t) \xrightarrow{d} MN \left(0, 2 \int_0^1 \sigma^4(t)dt \right). \quad (14)$$

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$$\sqrt{m} \left[RV^{(m)}(p) - IV \right] \xrightarrow{d} MN \left(0, \frac{4 - 2\pi}{\pi} \int_0^1 \sigma^4(t) dt \right), \quad (16)$$

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ML – General Results

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- Closed-form approximations of transition density (Aït-Sahalia, 2002, Aït-Sahalia and Yu, 2006).

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Exact and Approximate ML Estimation						
True Value $\kappa = 0.1$						
	Exact	Euler	Milst	Nowman	Infill	Herm
Mean	.2403	.2419	.2444	.2386	.2419	.2400
Std	.2777	.2867	.2867	.2771	.2867	.2762
RMSE	.3112	.3199	.3210	.3098	.3199	.3096
Mean of diff	NA	.0016	.0041	-.0017	.0016	-.0003
Std of diff	NA	.0500	.0453	.0162	.0500	.0043

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Parameter		κ	Bond Price	Option Price
True Value		0.1	0.8503	2.3920
Exact ML of CIR	Mean	0.1845	0.8438	1.8085
	RMSE	0.1319	0.0103	0.9052
Euler ML of CIR	Mean	0.1905	0.8433	1.7693
	RMSE	0.1397	0.0111	0.9668
ML of Vasicek (misspecified)	Mean	0.1746	0.8444	1.8837
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- Second source of bias – nonlinearity in pricing

Jackknife, Phillips and Yu, 2005, RFS

- Let N be the number of observations in the whole sample and decompose the sample into m consecutive subsamples each with ℓ observations, so that $N = m \times \ell$. Phillips and Yu (2005a) proposed the following jackknife estimator of κ , estimator

$$\hat{\kappa}_{jack} = \frac{m}{m-1} \hat{\kappa}_N - \frac{\sum_{i=1}^m \hat{\kappa}_{li}}{m^2 - m}, \quad (19)$$

Continue

- To reduce the bias in pricing, jackknife prices directly

$$\hat{P}_{jack} = 2\hat{P}_T - \frac{1}{2}(\hat{P}_{1,T/2} + \hat{P}_{2,T/2}).$$

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